



EUROPEAN CENTRAL BANK

EUROSYSTEM

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Assessing capital-based macroprudential policy using an integrated Early Warning (EW-) GVAR Model*

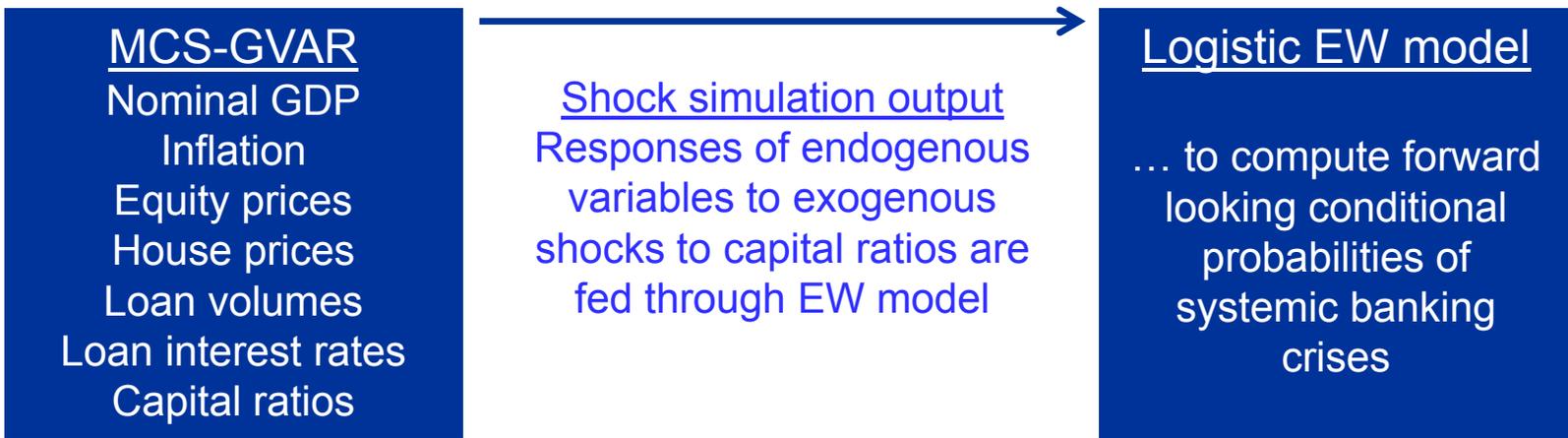
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* The views expressed are those of the authors and do not necessarily reflect those of the ECB or the ESRB.

Motivation

- We develop a model that can help assess **when capital-based macroprudential measures should be implemented + help calibrate them**
- Distinguish between costs and benefits of potential measures
 - **Benefit:** reduction in the probability of systemic banking sector crises
 - **Cost:** potential output loss due to implementation of measures
- Shock to capital ratios likely to come along with endogenous response of macro-financial variables included in logistic early warning models
- Integrate an **Early Warning (EW) Model** with a **Global Vector Autoregressive (GVAR) model**
- Why: to **endogenise predictor variables** contained in early warning model

The integrated Early Warning (EW-) GVAR Model



- Model accounts for **direct effects** of higher bank capitalization and for **feedback effects** from responses of macro-financial variables
- It helps shape our understanding of **transmission channels** of capital-based measures
- Can be used to analyze cross-country **spillover effects**
- Has some elements of **reciprocity** built-in

Contribution

- 1) Develop an integrated framework with **consistent multi-country macro / multi-banking system model** at its basis
 - many existing models based on area aggregates (euro area)
→ risk: **aggregation (attenuation) bias**
 - Explicit channels for **cross-border transmission** of capital ratio shocks, which existing models do not capture
- 2) Concerning **identification**: propose to derive bounds in terms of credit supply reaction
 - Distinguish between **how banks get to higher capital ratios**
 - Do not attempt to provide point estimates of reaction of lending and macro variables to capital ratio shock, but rather derive **bounds**
 - Do not make overly conservative assumptions about funding cost impact and pass-through to lending spreads

Overview

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| 1 | Motivation |
| 2 | Data |
| 3 | Model setup |
| 4 | Results |
| 5 | Conclusion |

Data

- Sample of 18 EU countries, 1980Q2-2014Q4
- **Dependent variable in EW model:**
 - Indicator of systemic financial stress events from the database recently created by Duprey, Klaus and Peltonen (2015)
 - Evaluation horizon: 7-12 quarters prior to the crisis
- **Independent variables in EW and basis for GVAR:**
 - BIS long series on domestic bank credit and total credit to the private non-financial sector
 - Macro-financial indicators: GDP, inflation, stock prices, house prices (from various sources, sourced through Haver Analytics)
 - Banking sector capitalization (BSI/OECD)
- **Output cost of a banking crisis:**
 - Average cost of a banking crisis obtained from Laeven and Valencia (2012)
 - Cumulative sum of the differences between actual and trend real GDP over the period of one year, expressed as a percentage of trend real GDP → output loss due to the crisis

Crisis dates and costs of banking crises

| | Crisis dates | Average output loss associated with a banking crisis (in % of GDP) |
|----------------|--|--|
| Austria | 2008q1-2010q3 | 14 |
| Belgium | 1990q3-1993q3, 2007q4-2013q1 | 19 |
| Germany | 1980q2-1982q1, 1992q3-1994q3, 2001q4-2003q3, 2008q3-2010q2 | 11 |
| Denmark | 1992q3-1993q3, 2008q1-2010q2 | 36 |
| Estonia | — | — |
| Spain | 1992q3-1992q3, 2008q1-ongoing | 49 |
| Finland | 1990q4-1996q2, 2001q1-2001q3, 2008q4-2010q3 | 70 |
| France | 1991q2-1993q1, 2008q1-2012q3 | 23 |
| Greece | 2008q1-ongoing | 43 |
| Ireland | 2008q1-ongoing | 106 |
| Italy | 1991q3-1996q3, 2008q1-ongoing | 32 |
| Luxembourg | 2008q2-ongoing | 36 |
| Netherlands | 1980q2-1983q3, 2002q2-2004q2, 2008q1-2010q3 | 23 |
| Portugal | 2008q1-ongoing | 37 |
| Sweden | 1982q1-1983q2, 1991q1-1994q3, 2000q4-2001q3, 2008q3-2010q3 | 29 |
| Slovenia | 2008q2-ongoing | 39 |
| Slovakia | — | — |
| United Kingdom | 2007q4-2010q2 | 25 |

Source: Duprey, Klaus, and Peltonen (2015), Laeven and Valencia (2012)

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Logistic regressions

$$P(y_{it} = 1) = \frac{e^{\alpha_i + X'_{it}\beta}}{1 + e^{\alpha_i + X'_{it}\beta}}$$

Note: The dependent variable is equal to one seven to twelve quarters preceding a banking crisis in a respective country and zero otherwise. Observations for crisis periods are omitted. Robust standard errors adjusted for clustering at the quarterly level are reported in parentheses. * indicates statistical significance at the 10 %-level, ** at the 5 %-level, and *** at the 1 %-level.

| Dependent variable | <i>Indicator for vulnerable state</i> | |
|--------------------------------|---------------------------------------|----------------------|
| | (1) | (2) |
| Credit growth | 0.142*** (0.030) | |
| Bank credit growth | | 0.256*** (0.035) |
| GDP growth | -0.109 (0.086) | -0.190* (0.101) |
| Inflation | -0.014 (0.095) | 0.063 (0.102) |
| House price growth | 0.184*** (0.023) | 0.183*** (0.024) |
| Equity price growth | 0.014** (0.007) | 0.013** (0.006) |
| Bank capitalization | -0.674*** (0.142) | -0.726*** (0.144) |
| Country dummies | YES | YES |
| Observations | 878 | 878 |
| Pseudo R-Squared | 0.216 | 0.279 |
| AUROC | 0.806 | 0.846 |
| Preference parameter (μ) | 0.85 | 0.85 |
| Optimal threshold | 64 | 71 |
| Type 1 errors | 0.255 | 0.276 |
| Type 2 errors | 0.293 | 0.214 |
| Relative usefulness | 0.421 | 0.477 |
| Adj noise-to-signal | 0.394 | 0.296 |
| Perc predicted | 0.745 | 0.724 |
| Cond prob | 0.334 | 0.401 |
| Prob diff | 0.169 | 0.236 |

Mixed cross-section GVAR model

- Adopted from Gross, Kok, and Zochowski (2015)
- Mixed cross-section GVAR with two equation blocks: countries and banking systems; see [Annex for details](#)
- Different weights to link different cross sections (help reflect possible transmission channels)
- Note: [For a bank it doesn't matter how much the country in which it's located trades with other countries. Its own exposure matters instead.](#)

| From... | To... | Banking system-based model |
|------------------|------------------|---|
| Countries - | - Countries | Bilateral trade (sum of nominal inports and exports) |
| | - Banking system | Transpose of Banking system (banks) - Countries matrices |
| Banking system - | - Countries | BSI domestic and cross-border exposure data |
| | - Banking system | BSI cross-banking system exposure to financial institutions |

Defining deleveraging

- Deleveraging means that the leverage ratio (TA/E) comes down, i.e. capital ratio up; can be achieved via one or combination of:

$A\downarrow \quad E\rightarrow \quad D\downarrow$: reduce assets via selling off business, disposal of liquid assets, non-renewal of maturing loans

$A\uparrow \quad E\uparrow \quad D\rightarrow$: raise equity capital and invest in new assets

$A\rightarrow \quad E\uparrow \quad D\downarrow$: raise equity and replace debt (i.e. constant BS)

- Retaining earnings to gradually build up E falls into 2nd (or 3rd) category
- First two reactions form starting point for shock simulations with the integrated EW-GVAR model

Shock identification – Simulation types

Scenario Type 1: Banks get to higher capital ratio (Δ) by letting business mature or selling assets \rightarrow negative credit supply shock (“**lower bound**”)

\rightarrow Sign constraints: loans down, loan interest rates up

Scenario Type 2: Banks get to higher capital ratio (Δ) by raising capital and investing \rightarrow positive credit supply shock (“**upper bound**”)

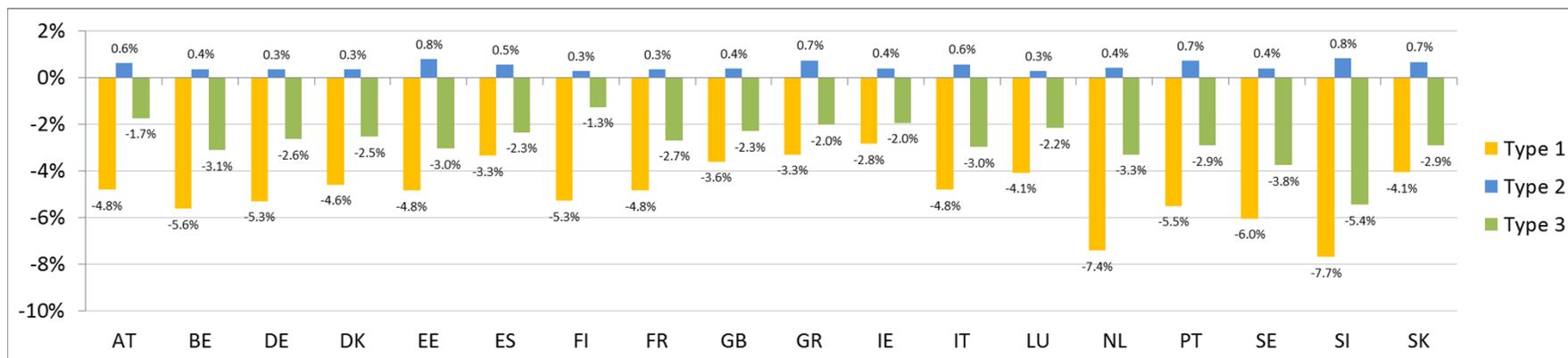
\rightarrow Sign constraints: loans up, loan interest rates down

Scenario Type 3: Unconstrained capital ratio shock (Δ) to see how banks went over deleveraging process historically

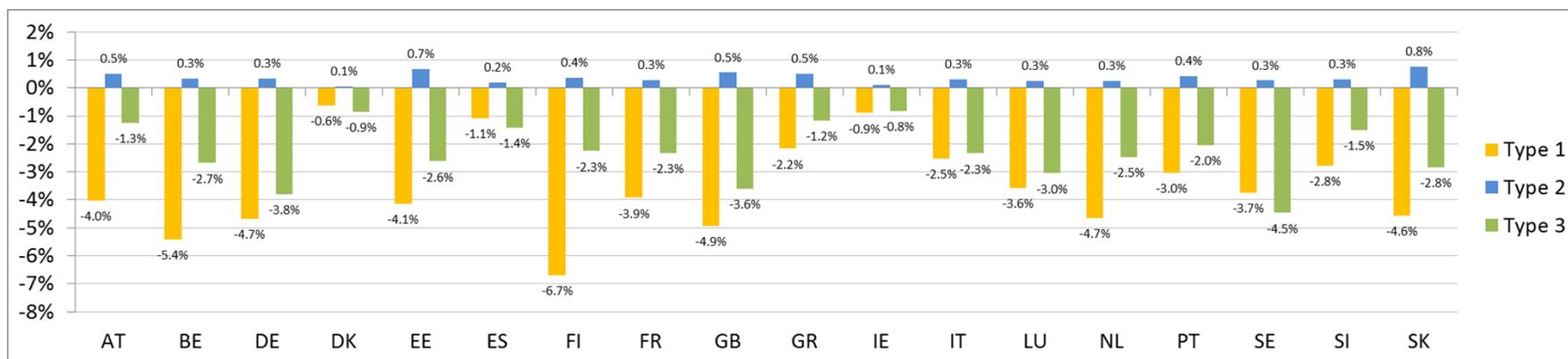
\rightarrow No sign constraints

Simulation results – Nominal credit and real GDP

Cumulative loan growth effects under Type 1/2/3 (100 bps capital ratio shock):



Cumulative real GDP growth responses under Type 1/2/3:



Feeding simulated shocks/responses through EWM

- Feed simulated shocks and responses for 2015Q1-2017Q4 through logistic model to obtain predicted probabilities of being in vulnerable state
- **Change in the probability of being in a vulnerable state:** difference between average predicted probability over simulation horizon and predicted probability in 2014Q4:

$$\Delta p = \overline{p_{sim}} - p_{2014Q4}$$

- **Benefit of capital ratio shocks:** product of reduction in crisis probabilities and average output loss associated with banking sector crises in respective country (output loss estimate from Laeven and Valencia 2012):

$$benefit = -\Delta p * crisis\ cost$$

The net benefit of capital-based measures

- *Net benefit* of capital-based measures:

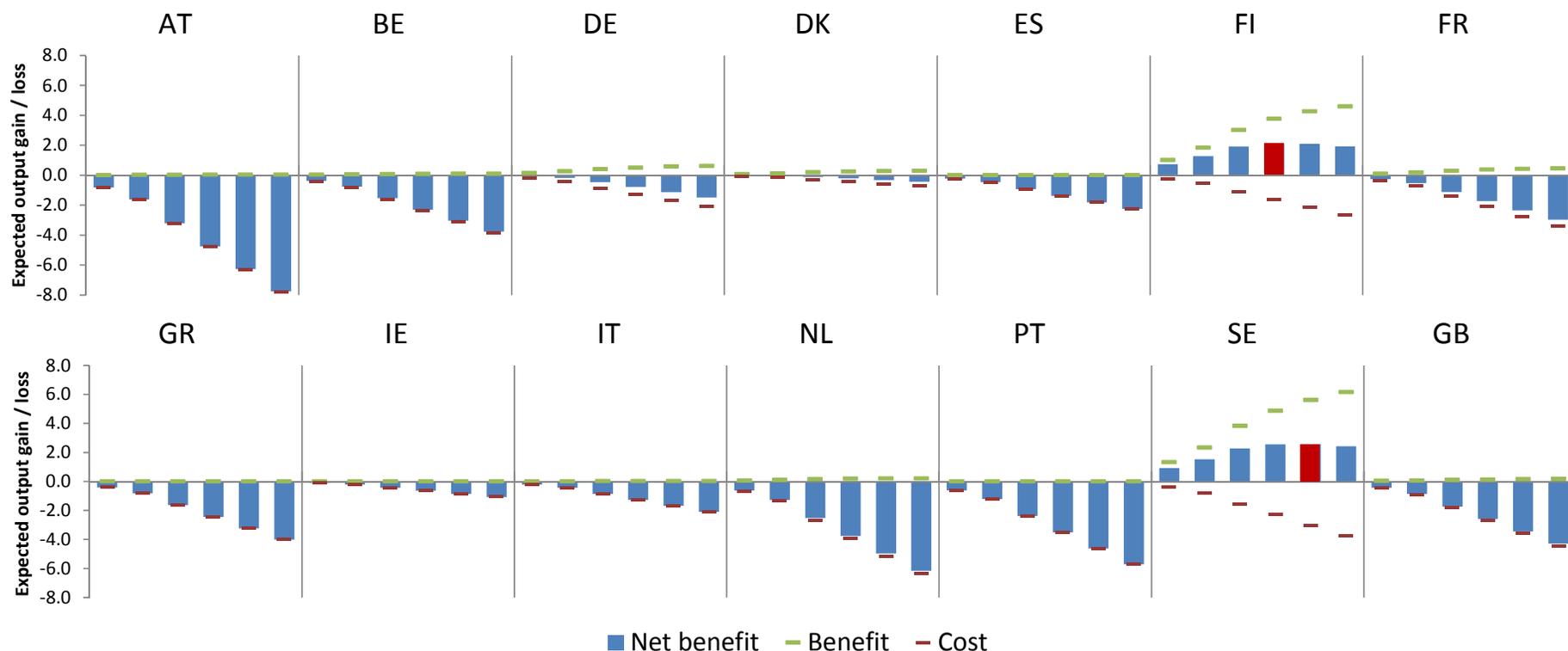
$$\textit{net benefit} = \textit{benefit} - \textit{cost}$$

- *Cost* estimate from GVAR: cumulative output loss relative to baseline
- *Net benefit* can be interpreted as **expected output gain or loss** from imposing capital ratio shock on banking system
- Assess the *net benefit* for the 3 types of shocks
 - Type 1: contractionary deleveraging
 - Type 2: expansionary deleveraging
 - Type 3: unconstrained deleveraging

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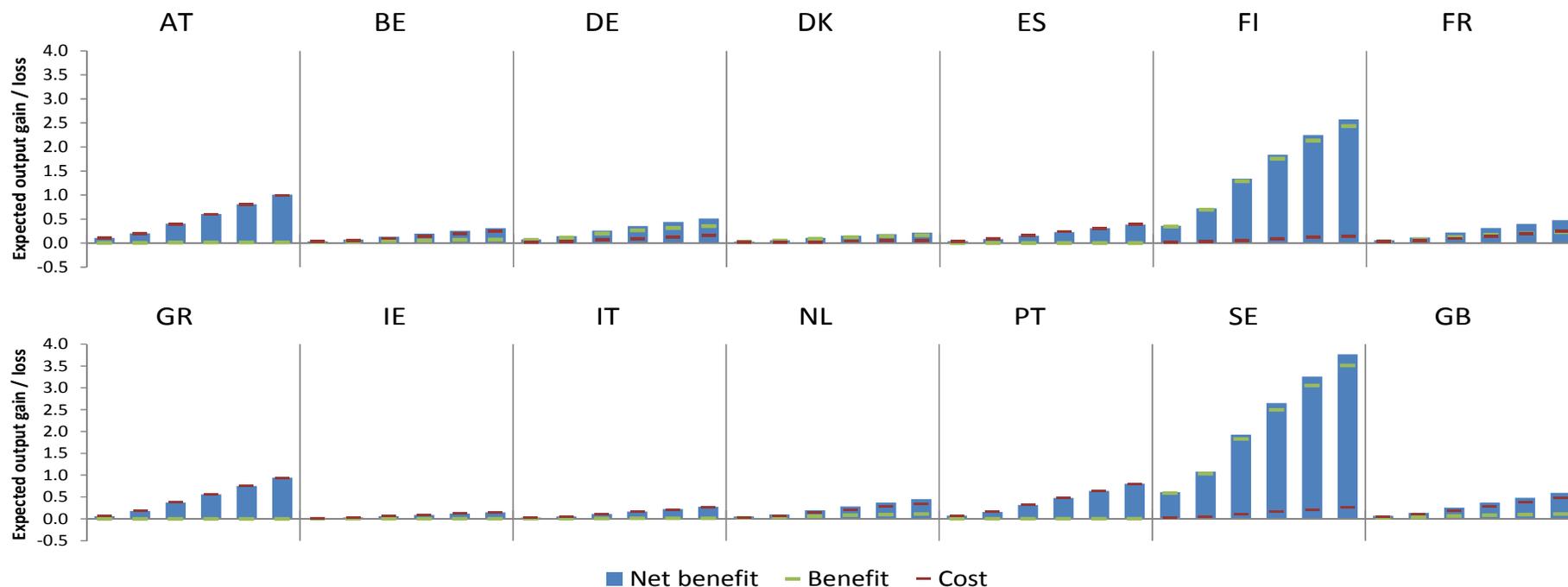
Net benefits – contractionary deleveraging (Type 1)



Note: Bars and markers for each country refer to capital ratio shocks of 25, 50, 100, 150, 200, and 250 bps, respectively.

- Cross-country heterogeneity: countries with predicted probabilities close to 0 in 2014Q4 do not have much to gain from capital-based measures
- Only FI and SE would benefit at the current stage of the cycle

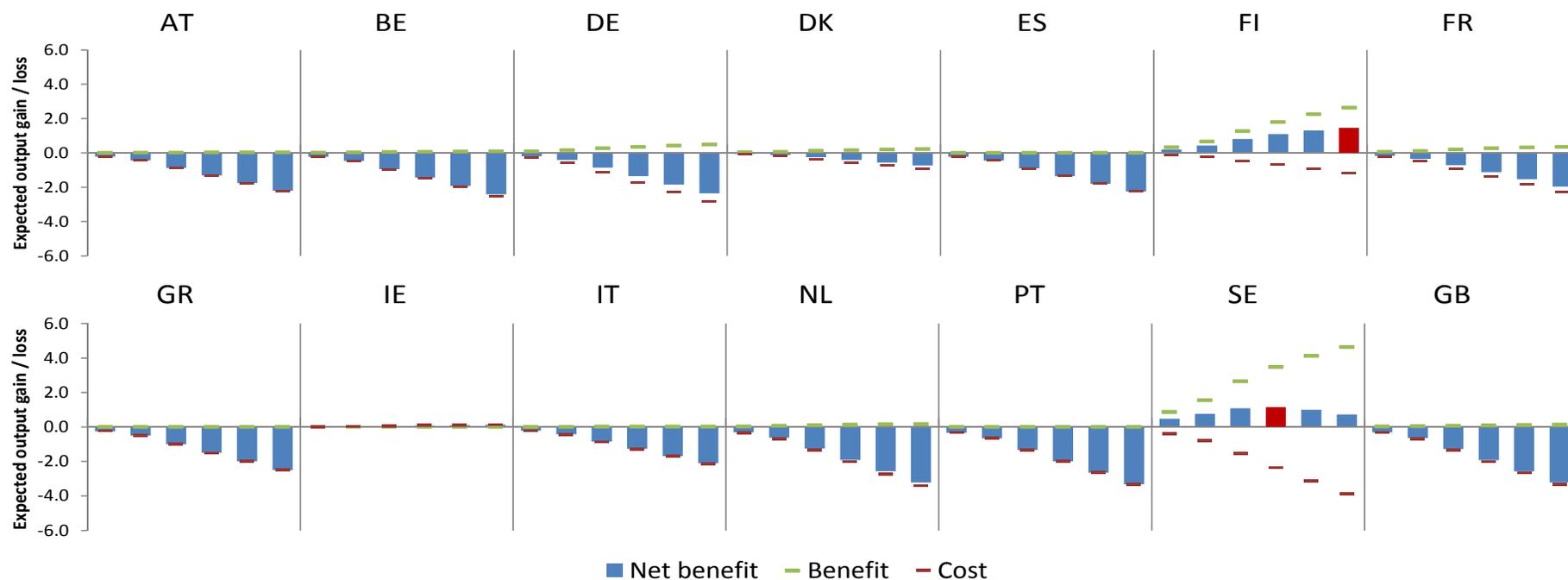
Net benefits – expansionary deleveraging (Type 2)



Note: Bars and markers for each country refer to capital ratio shocks of 25, 50, 100, 150, 200, and 250 bps, respectively.

- Predicted probabilities decrease, since the effects of higher bank capitalization dominate the effects of (slightly) increased credit growth
- GDP reacts positively to the positive credit supply shock → no costs

Net benefits – unconstrained deleveraging (Type 3)



Note: Bars and markers for each country refer to capital ratio shocks of 25, 50, 100, 150, 200, and 250 bps, respectively.

- Results for Type 3 scenario fall in between Type 1 and 2
- Closer to type 1, since credit/asset prices & GDP tend to react negatively

Interpretation and further results

- Benefits depend on **relative importance of individual variable reactions**:
 - higher credit / asset price growth $\Rightarrow p \uparrow$, benefit \downarrow
 - higher GDP growth / bank capitalization $\Rightarrow p \downarrow$, benefit \uparrow } different reactions across scenario types
- Model allows to analyse **contribution of individual variables** on Δp :
 - **Type 1 & 3**: $p \downarrow$ mainly driven by higher bank capitalization and lower credit growth; to a lesser extent: lower house price growth; countervailing: lower GDP growth
 - **Type 2**: $p \downarrow$ almost entirely driven by higher bank capitalization; countervailing effects due to higher credit growth are of minor importance
- Model allows to analyse **spillover effects** since GVAR model generates domestic and foreign responses
 - two channels: (1) cross-border lending; and (2) trade
 - significant (negative) spillover effects particularly for larger countries
 - aggregate foreign effects tend to go on in same direction as domestic effects

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Conclusions

- Integrated tool for informing **activation** and **calibration** of capital-based measures
- Account for **costs** and **benefits** of capital-based measures
- Results (in particular short-run macro costs) **depend on how banks move to higher capital ratios**
- Effects of dampened credit/asset price growth can be sizable: under Type 1 shocks they can account for >50% of the reduction in predicted probabilities
 - macroprudential policy should be ambitious in **smoothing the cycle**
- Not much potential for countercyclical measures at the current stage of the cycle (exception: Finland and Sweden)
- Spillover effects mostly in larger countries, under Type 1 and 3 shocks

Conclusions (ctd)

- Additional steps:
 - Robustness checks with respect to prediction horizon (currently 7-12 quarters before crisis), crisis definition, and out-of-sample performance
 - For benefit calculations:
 - Estimate average output loss during crises based on our own data, instead of referring to Laeven / Valencia (2012)
 - Use cross-country average (or other moments of the distribution) of GDP losses during crises instead of country specific crisis costs
 - Use alternative evaluation method for the benefits that makes stronger use of the model's binary signalling properties (signal vs. no signal)



Background slides

Related literature

- **Early warning literature:** Kaminsky et al. (1998), Alessi and Detken (2011), Behn et al. (2013), Lo Duca and Peltonen (2013), Betz et al. (2014)
- **Effects of macroprudential instruments:** Lim et al. (2011), Galati and Moessner (2014), Bruno et al. (2015), Cerutti et al. (2015)
- **Relationship between capital ratios and loan growth:** Bernanke and Lown (1991), Gambacorta and Mistrulli (2004), Francis and Osborne (2009), Berrospide and Edge (2010), Bochmann et al. (2015), Gross et al. (2015)
- **Empirical literature on capital requirements and lending:** Brun et al. (2013), Carlson et al. (2013), Aiyar et al. (2014), Jimenez et al. (2014), Behn et al. (2015), Marques-Ibanez et al. (2015)
- See Annex to the slides and draft paper for more references

Literature on impact of higher capital requirements

- Assumed trade-off between financial stability and cost of financial intermediation
- Concave function between capital and welfare implies an optimal level of capital exists
- Operating below optimal capital level means lower financial intermediation cost in short run but higher probability of bank or banking system default (costly)
- Operating above optimal capital level means too high financial intermediation cost, more than outweighing the gain from lower crisis probability
- **BIS 2010**: net benefit from increasing cap ratios positive for broad range
- **Miles/Yang/Marcheggiano 2012**: optimal bank capital should be around 20%
- **Kashyap et al. 2010**: focus on cost side; though conclude with respect to net benefit: likely positive because lending rates will not react much

Literature on impact of higher capital requirements

- Higher capital induces banks to better screen borrowers (Coval and Thakor 2005) and to more efficiently monitor them (Holmstrom and Tirole 1997 and Mehran and Thakor 2011)
- Based on Miller 1995 and the Modigliani-Miller capital structure irrelevance theorem: Admati et al. (2011) and Admati and Hellwig (2013) call for much higher capital ratios than currently imposed; social costs of significantly higher requirements will be negligible and the benefit of reduced probabilities of failure far outweigh these costs
- Further references: Beguenau (2014), Nguyen (2014), van den Heuvel (2008), Calomiris and Herring (2013)

Micro-econometric studies to measure impact of changes in capital requirements

- Micro (bank-level) data helps for shock identification
- Effects on an increase in capital requirements on bank lending

In the transition...

| | pp-variation in cap. Req. | Effect on bank lending volume | Horizon |
|----------------------------|---------------------------|-------------------------------|-------------|
| Maurin and Toivanen (2012) | +1pp | -2.15% | medium term |
| Aiyar et al. (2014) | +1pp | -8.40% | 1 year |
| Noss and toffano (2014) | +1pp | -4.50% | 3 years |
| Mesonnier and Monks (2014) | +1pp | -1.40% | 1 year |
| Brun et al. (2015) | -1pp | 5.00% | short term |

... and in the long run

| | pp-variation in cap. Req. | Effect on bank lending rate |
|-------------------------|---------------------------|-----------------------------|
| Kashyap et al. (2010) | +1pp | +0.03pp |
| King (2010) | +1pp | +0.15pp |
| Kisin and Manela (2015) | +1pp | +0.003pp |

Source: From the response to the EC Consultation on the impact of CRR/CRD4 on bank lending, prepared jointly with I. Jaccard, F. Boissay

Recognition provisions in EU law

| Macro-prudential measure | Legal basis | Recognition in EU law |
|--------------------------------|--------------------|-------------------------|
| Countercyclical capital buffer | 130, 135-140 CRDIV | Mandatory (up to 2.5 %) |
| Higher RWs for SA banks | 124 CRR | Mandatory |
| Higher LGDs for IRB banks | 164 CRR | Mandatory |
| National flexibility measures | 458 CRR | Voluntary |
| Systemic risk buffer | 133-134 CRDIV | Voluntary |
| Pillar II measures | 103 CRDIV | Not mentioned |
| G-SII buffer | 131 CRDIV | Not mentioned |
| O-SII buffer | 131 CRDIV | Not mentioned |

Source: IWG Expert Group on cross-border effects of macro-prudential policy and reciprocity, Report to the ATC.

Transmission channels for capital-based policy

- Broad capital-based instrument applied to bank in Country A which does significant x-border business in Country B likely causes **x-border loan supply** change in Country B
- Even if banks are active only in home country, change in their capital and hence domestic loan supply may feed first through domestic macro but then to Country B through the **trade channel**
- **Both channels well captured in the MCS-GVAR**
- Distinguish between x-border business through subsidiaries vs branches vs direct cross-border lending
- BSI data: carving out only direct x-border lending which we add to banking system credit; meaningful because direct x-border lending likely among the most reactive to capital change for parent bank

Shock identification

- Observed capital ratios very much endogenous
- Capital ratio change as such should not have any macro impact as long as prices or volumes (loan supply) are not affected
- We circumvent identification based on capital by first translating capital ratio shocks into **credit supply shocks, considering two polar cases of reaction (seen as upper and lower bound reaction)**
- Identification based on **sign restrictions** for credit supply shocks

Shock identification – Simulation types

Scenario Type 1: Banks get to higher capital ratio (Δ) by letting business mature or selling assets \rightarrow negative credit supply shock (“**lower bound**”)

$$shock^{Type1} = \ln\left(\frac{E_0}{E_0/A_0 + \Delta}\right) - \ln(A_0) \quad [A_0: \text{assets at } T_0, E_0: \text{equity at } T_0, \Delta: \text{capital ratio shock in p.p.}]$$

Scenario Type 2: Banks get to higher capital ratio (Δ) by raising capital and investing \rightarrow positive credit supply shock (“**upper bound**”)

$$shock^{Type2} = \ln\left(A_0 - E_0 \left(\frac{(\Delta + E_0/A_0)(A_0 - E_0)}{E_0(\Delta + E_0/A_0 - 1)} + 1\right)\right) - \ln(A_0)$$

Scenario Type 3: Unconstrained capital ratio shock (Δ) to see how banks went over deleveraging process historically

| # | Scenario | Shock | Sign constraints |
|---|---|----------|------------------|
| 1 | Contractionary deleveraging shock | LEV down | L down, I up |
| 2 | Expansionary deleveraging shock | LEV down | L up, I down |
| 3 | Deleveraging shock -- No sign constraints | LEV down | - |

The semi-structural MCS-GVAR model

The equation system

$$x_{it} = a_i + \sum_{p_1=1}^{P_1} \Phi_{ip_1} x_{i,t-p_1} + \sum_{p_2=0}^{P_2} \Lambda_{i,0,p_2} x_{i,t-p_2}^{*,C-C} + \sum_{p_3=0}^{P_3} \Lambda_{i,1,p_3} y_{i,t-p_3}^{*,C-B} + \sum_{p_4=0}^{P_4} \Lambda_{i,2,p_4} z_{i,t-p_4}^{*,C-CB} + \sum_{p_5=0}^{P_5} K_{i,p_5} v_{t-p_5} + \varepsilon_{it}$$

$$y_{jt} = b_j + \sum_{q_1=1}^{Q_1} \Pi_{jq_1} y_{j,t-q_1} + \sum_{q_2=0}^{Q_2} \Xi_{j,0,q_2} x_{j,t-q_2}^{*,B-C} + \sum_{q_3=0}^{Q_3} \Xi_{j,1,q_3} y_{j,t-q_3}^{*,B-B} + \sum_{q_4=0}^{Q_4} \Xi_{j,2,q_4} z_{j,t-q_4}^{*,B-CB} + \sum_{q_5=0}^{Q_5} E_{j,q_5} v_{t-q_5} + \omega_{jt}$$

$$z_{lt} = c_l + \sum_{r_1=1}^{R_1} \Gamma_{lr_1} z_{l,t-r_1} + \sum_{r_2=0}^{R_2} \Psi_{l,0,r_2} x_{l,t-r_2}^{*,CB-C} + \sum_{r_3=0}^{R_3} \Psi_{l,1,r_3} y_{l,t-r_3}^{*,CB-B} + \sum_{r_4=0}^{R_4} \Psi_{l,2,r_4} z_{l,t-r_4}^{*,CB-CB} + \sum_{r_5=0}^{R_5} T_{l,r_5} v_{t-r_5} + \tau_{lt}$$

Three sets of equations for **countries**, banks, and **central banks**.

Global exogenous (v_t) or local exogenous variables can be included. Global exogenous for instance oil price. Local exogenous can be anything, depending on needs and type of application/simulation.

The semi-structural MCS-GVAR model

The equation system

$$\begin{aligned}
 x_{it} &= a_i + \sum_{p_1=1}^{P_1} \Phi_{ip_1} x_{i,t-p_1} + \sum_{p_2=0}^{P_2} \Lambda_{i,0,p_2} x_{i,t-p_2}^{*,C-C} + \sum_{p_3=0}^{P_3} \Lambda_{i,1,p_3} y_{i,t-p_3}^{*,C-B} + \sum_{p_4=0}^{P_4} \Lambda_{i,2,p_4} z_{i,t-p_4}^{*,C-CB} + \sum_{p_5=0}^{P_5} K_{i,p_5} v_{t-p_5} + \varepsilon_{it} \\
 y_{jt} &= b_j + \sum_{q_1=1}^{Q_1} \Pi_{jq_1} y_{j,t-q_1} + \sum_{q_2=0}^{Q_2} \Xi_{j,0,q_2} x_{j,t-q_2}^{*,B-C} + \sum_{q_3=0}^{Q_3} \Xi_{j,1,q_3} y_{j,t-q_3}^{*,B-B} + \sum_{q_4=0}^{Q_4} \Xi_{j,2,q_4} z_{j,t-q_4}^{*,B-CB} + \sum_{q_5=0}^{Q_5} P_{j,q_5} v_{t-q_5} + \omega_{jt} \\
 z_{lt} &= c_l + \sum_{r_1=1}^{R_1} \Gamma_{lr_1} z_{l,t-r_1} + \sum_{r_2=0}^{R_2} \Psi_{l,0,r_2} x_{l,t-r_2}^{*,CB-C} + \sum_{r_3=0}^{R_3} \Psi_{l,1,r_3} y_{l,t-r_3}^{*,CB-B} + \sum_{r_4=0}^{R_4} \Psi_{l,2,r_4} z_{l,t-r_4}^{*,CB-CB} + \sum_{r_5=0}^{R_5} T_{l,r_5} v_{t-r_5} + \tau_{lt}
 \end{aligned}$$

Fully endogenous (though constrained) cross-cross-section dependence via weighted variable vectors (“star-variables”).

To establish link b/w 3 cross-sections, up to 9 weight matrices needed. Some weight sets not needed due to exclusion restrictions.

The semi-structural MCS-GVAR model

The equation system

- See the difference between MCS structure and standard GVAR with variable-specific weights!
- Example: loan growth equation
 - Standard GVAR w/ v-specific weights: e.g. BIS weights on credit but trade weights on GDP (!)
 - MCS-GVAR: exposure of bank or banking system vis-à-vis countries, i.e. reflecting its activity there and hence susceptibility to macro
- For a bank it doesn't matter how much the country in which it's located trades with other countries; its x-border exposure matters
- Predictive performance tests confirm MCS makes more sense

The semi-structural MCS-GVAR model

Solving the global model – STEP 1/4

Step 1: Generate A-matrices. One starts by stacking the within-cross-section vectors along with the cross-cross-section weighted variable vectors in (here) three vectors m_{it}^x , m_{jt}^y , and m_{lt}^z .

$$\begin{aligned} m_{it}^x &= \left(x_{it} \quad x_{it}^{*,C-C'} \quad y_{it}^{*,C-B'} \quad z_{it}^{*,C-CB'} \right)' \\ m_{jt}^y &= \left(y_{jt} \quad x_{jt}^{*,B-C'} \quad y_{jt}^{*,B-B'} \quad z_{jt}^{*,B-CB'} \right)' \\ m_{lt}^z &= \left(z_{lt} \quad x_{lt}^{*,CB-C'} \quad y_{lt}^{*,CB-B'} \quad z_{lt}^{*,CB-CB'} \right)' \end{aligned}$$

The equation system can be re-written with these m vectors as follows.

$$\begin{aligned} \underbrace{\left(I_{k_i^x} \quad -\Lambda_{i,0,0} \quad -\Lambda_{i,1,0} \quad -\Lambda_{i,2,0} \right)}_{\equiv A_{i0}^x} m_{it}^x &= \mathbf{a}_i + \underbrace{\left(\Phi_{i1} \quad \Lambda_{i,1,1} \quad \Lambda_{i,2,1} \right)}_{\equiv A_{i1}^x} m_{i,t-1}^x + \dots + \epsilon_{it} \\ \underbrace{\left(I_{g_j^y} \quad -\Xi_{j,0,0} \quad -\Xi_{j,1,0} \quad -\Xi_{j,2,0} \right)}_{\equiv A_{j0}^y} m_{jt}^y &= \mathbf{b}_j + \underbrace{\left(\Pi_{j1} \quad \Xi_{j,1,1} \quad \Xi_{j,2,1} \right)}_{\equiv A_{j1}^y} m_{j,t-1}^y + \dots + \omega_{jt} \\ \underbrace{\left(I_{k_l^z} \quad -\Psi_{l,0,0} \quad -\Psi_{l,1,0} \quad -\Psi_{l,2,0} \right)}_{\equiv A_{l0}^z} m_{lt}^z &= \mathbf{c}_l + \underbrace{\left(\Gamma_{l1} \quad \Psi_{l,1,1} \quad \Psi_{l,2,1} \right)}_{\equiv A_{l1}^z} m_{l,t-1}^z + \dots + \tau_{lt} \end{aligned}$$

The semi-structural MCS-GVAR model

Solving the global model – STEP 2/4

Step 2: Generate L-matrices (“link” matrices). With a global, stacked variable vector $s_t = (x'_{1t}, \dots, x'_{Nt}, y'_{1t}, \dots, y'_{Mt}, z'_{1t}, \dots, z'_{Bt})$ at hand, the cross-section-specific variable vectors m^x_{it} , m^y_{jt} , and m^z_{lt} to s_t can be linked. The link matrices L^x_i , L^y_j , and L^z_l are used to map the local cross-section variables into the global vector, which involve the weights from the weight matrices W .

$$\begin{aligned} m^x_{it} = L^x_i s_t &\quad \rightarrow \quad A^x_{i0} L^x_i s_t = a_i + A^x_{i1} L^x_i s_{t-1} + \dots + \epsilon_{it} \\ m^y_{jt} = L^y_j s_t &\quad \rightarrow \quad A^y_{j0} L^y_j s_t = b_j + A^y_{j1} L^y_j s_{t-1} + \dots + \omega_{jt} \\ m^z_{lt} = L^z_l s_t &\quad \rightarrow \quad A^z_{l0} L^z_l s_t = c_l + A^z_{l1} L^z_l s_{t-1} + \dots + \tau_{lt} \end{aligned}$$

The semi-structural MCS-GVAR model

Solving the global model – STEP 3/4

Step 3: Generate G-matrices. The equation-by-equation system can now be stacked into a global system.

$$\begin{aligned} G_0^x &= \begin{pmatrix} A_{10}^x L_1^x \\ \dots \\ A_{N0}^x L_N^x \end{pmatrix}, G_1^x = \begin{pmatrix} A_{11}^x L_1^x \\ \dots \\ A_{N1}^x L_N^x \end{pmatrix}, \dots, a = \begin{pmatrix} a_1 \\ \dots \\ a_N \end{pmatrix} \\ G_0^y &= \begin{pmatrix} A_{10}^y L_1^y \\ \dots \\ A_{M0}^y L_M^y \end{pmatrix}, G_1^y = \begin{pmatrix} A_{11}^y L_1^y \\ \dots \\ A_{M1}^y L_M^y \end{pmatrix}, \dots, b = \begin{pmatrix} b_1 \\ \dots \\ b_M \end{pmatrix} \\ G_0^z &= \begin{pmatrix} A_{10}^z L_1^z \\ \dots \\ A_{B0}^z L_B^z \end{pmatrix}, G_1^z = \begin{pmatrix} A_{11}^z L_1^z \\ \dots \\ A_{B1}^z L_B^z \end{pmatrix}, \dots, c = \begin{pmatrix} c_1 \\ \dots \\ c_B \end{pmatrix} \end{aligned}$$

These cross-section-specific G matrices can be further combined to a set of global G matrices. The intercept vectors a , b , and c will be combined in a vector d . That is,

The semi-structural MCS-GVAR model

Solving the global model – STEP 4/4

Step 4: Generate H-matrices. The global system can now be pre-multiplied by the inverse of G_0 . The system is now ready to be used for shock simulation and forecast purposes.

$$s_t = \underbrace{G_0^{-1}d}_{\equiv H_0} + \underbrace{G_0^{-1}G_1}_{\equiv H_1}s_{t-1} + \dots + G_0^{-1}\varphi_t$$

Signalling approach

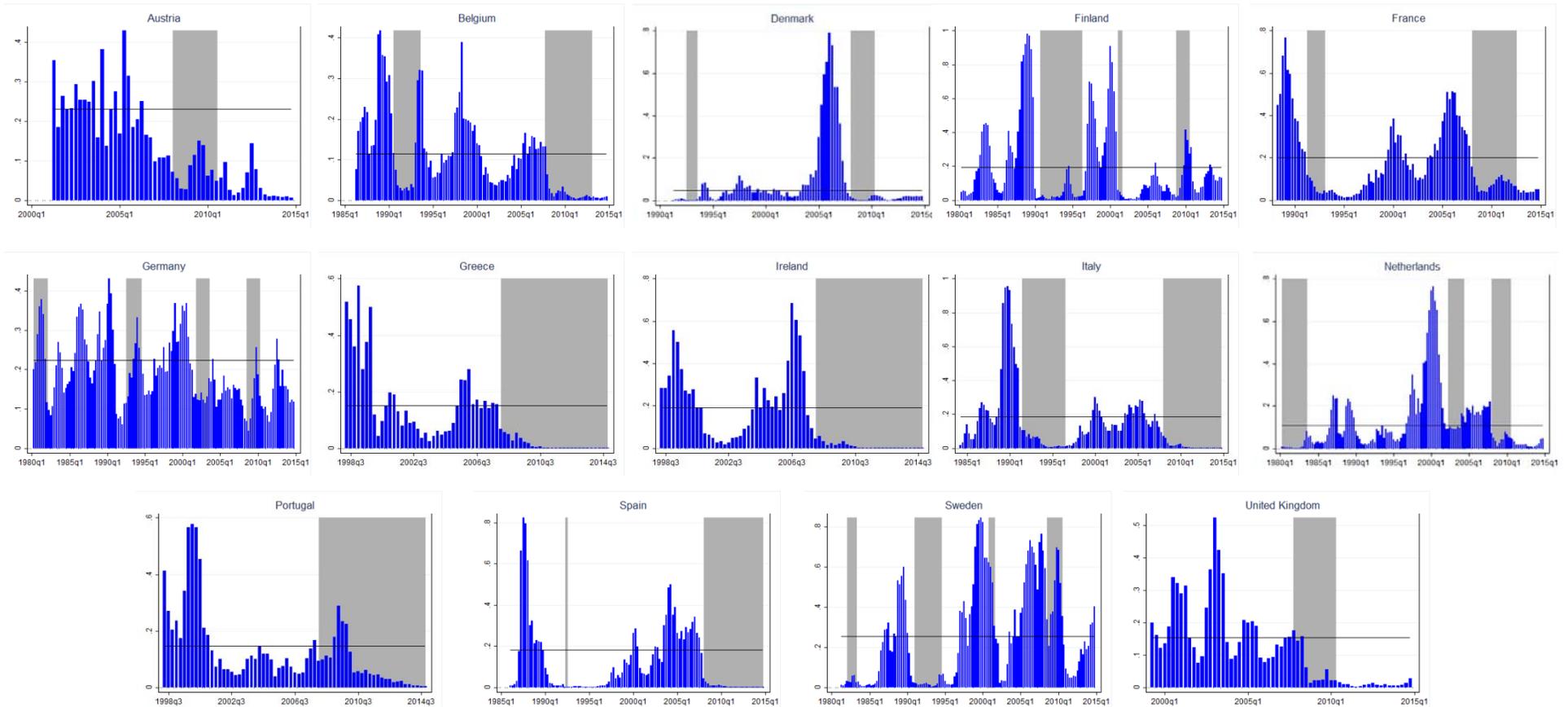
- Assume that a model issues a warning whenever the predicted probability surpasses a threshold τ , defined as a percentile of the country-specific distribution of predicted probabilities
- Compare signals issued by model to actual outcomes for each observation
- Optimal warning threshold depends on policy makers' relative aversion with respect to **type I ($T_1(\tau)$, missing crisis)** and **type II ($T_2(\tau)$, false alarm) errors**, and the unconditional probability of crisis (P_1) vs. non-crisis periods (P_2)
- **Optimal warning threshold minimizes** the following **loss function**:

$$L(\tau) = \mu * P_1 * T_1(\tau) + (1 - \mu) * P_2 * T_2(\tau)$$

- Benchmark: $\mu = 0.85$, policy maker is more averse against missing a crisis
- Absolute usefulness: $U_a = \min(\mu P_1, (1 - \mu) P_2) - L(\mu)$
- Relative usefulness: $U_r = U_a / \min(\mu P_1, (1 - \mu) P_2)$

Predicted probabilities of being in a vulnerable state

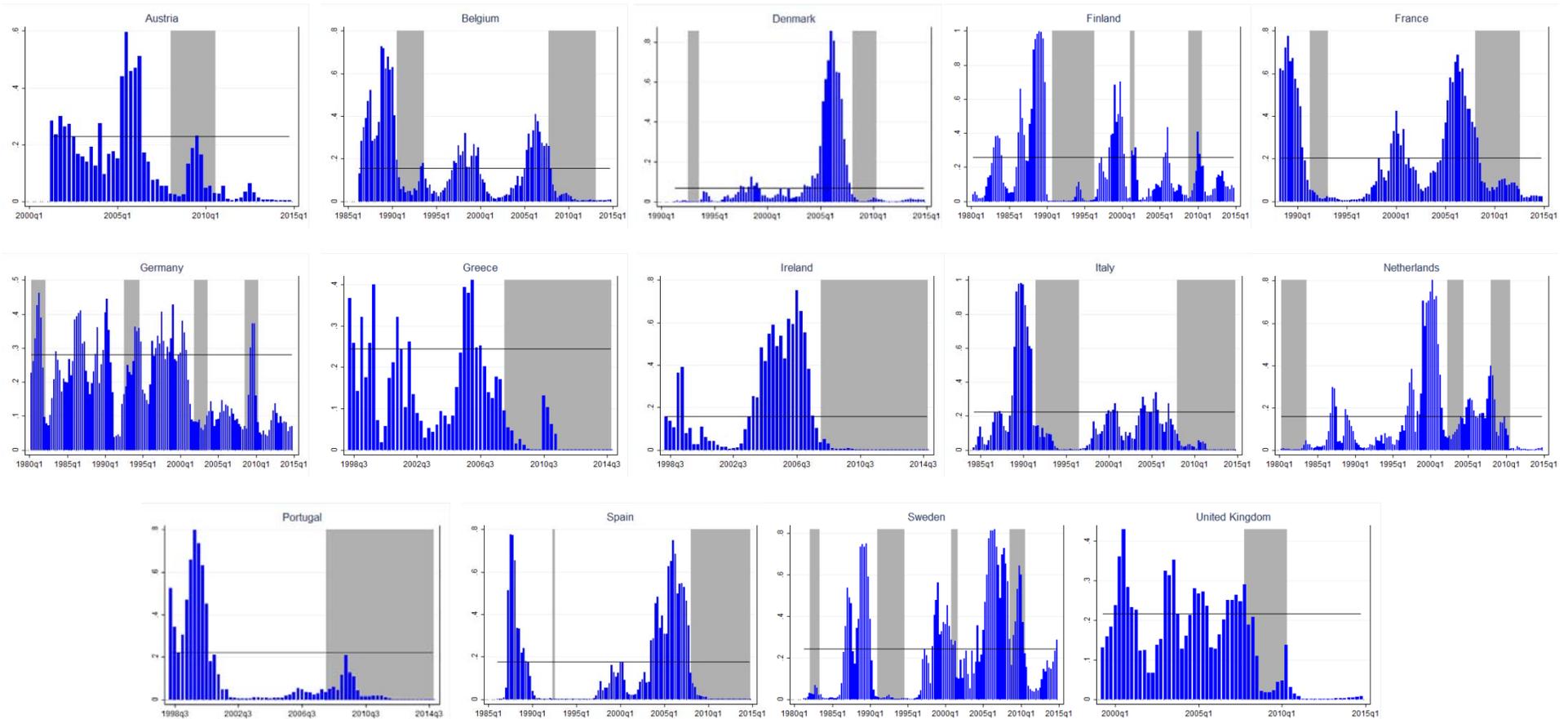
Model 4



Note: Predicted probabilities are depicted in blue. Horizontal black lines indicate the optimal warning threshold ($\mu=0.85$). Shaded grey areas correspond to systemic events.

Predicted probabilities of being in a vulnerable state

Model 8



Note: Predicted probabilities are depicted in blue. Horizontal black lines indicate the optimal warning threshold ($\mu=0.85$). Shaded grey areas correspond to systemic events.

Recalling the transmission channels

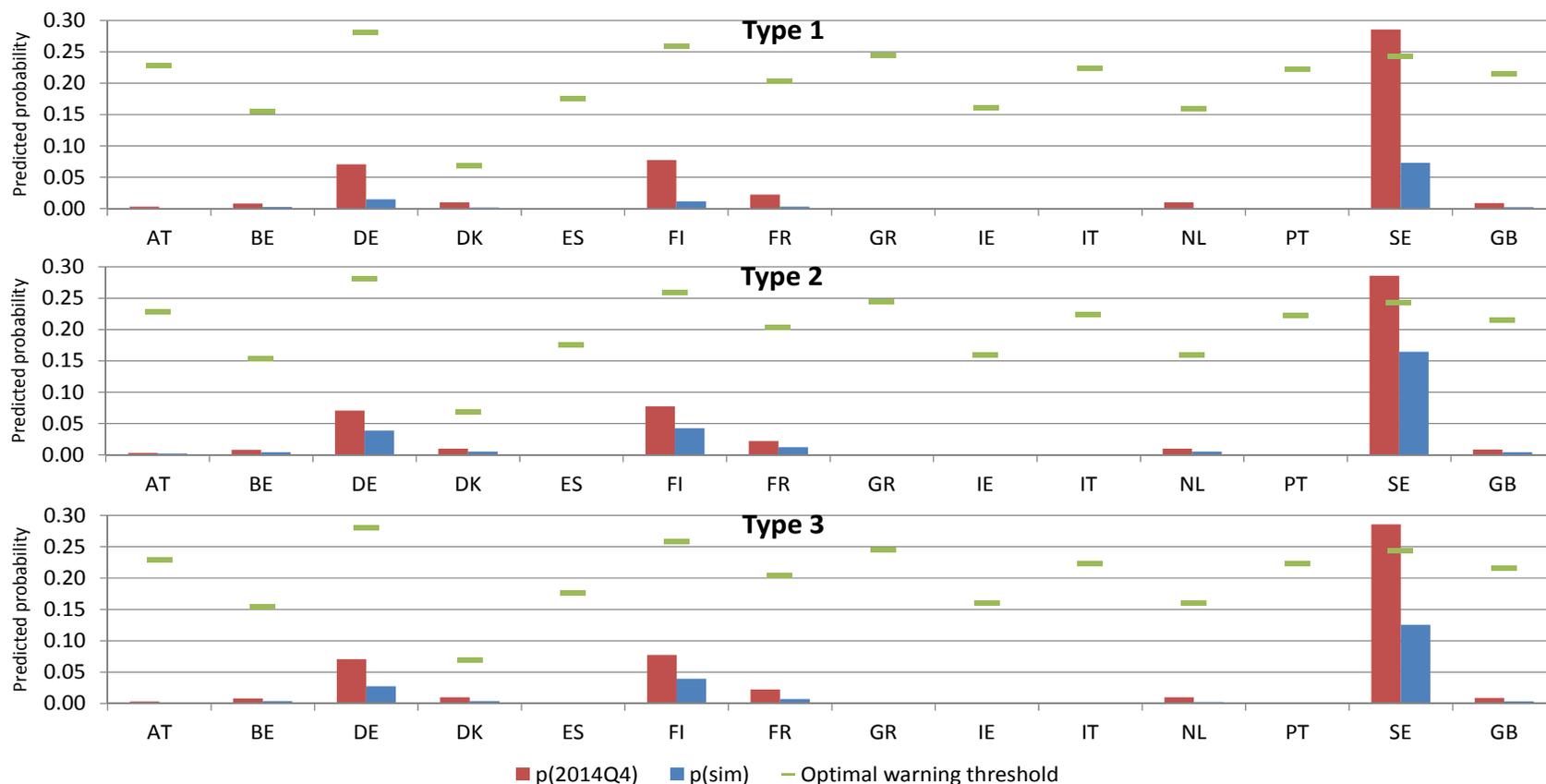
| | |
|---------------------|----------------------|
| Bank credit growth | 0.256*** (0.035) |
| GDP growth | -0.190* (0.101) |
| Inflation | 0.063 (0.102) |
| House price growth | 0.183*** (0.024) |
| Equity price growth | 0.013** (0.006) |
| Bank capitalization | -0.726*** (0.144) |

- Recall coefficients from logit model:
 - higher credit / asset price growth $\Rightarrow p \uparrow$
 - higher GDP growth / bank capitalization $\Rightarrow p \downarrow$

| Type 1 | Type 2 |
|---|--|
| • Credit growth $\downarrow \Rightarrow p \downarrow$ | • Credit growth $\uparrow \Rightarrow p \uparrow$ |
| • Asset prices $\downarrow \Rightarrow p \downarrow$ | • Asset prices $\uparrow \Rightarrow p \uparrow$ |
| • GDP growth $\downarrow \Rightarrow p \uparrow$ | • GDP growth $\uparrow \Rightarrow p \downarrow$ |
| • Bank capital $\uparrow \Rightarrow p \downarrow$ | • Bank capital $\uparrow \Rightarrow p \downarrow$ |

- Whether crisis probabilities go up or down depends on the relative importance of the effects \Rightarrow ‘benefit’ can be positive or negative
- Costs are estimated as the cumulative output loss or gain relative to the baseline from the GVAR \Rightarrow ‘cost’ positive under Type 2, negative under Type 1, and tends to be negative under Type 3

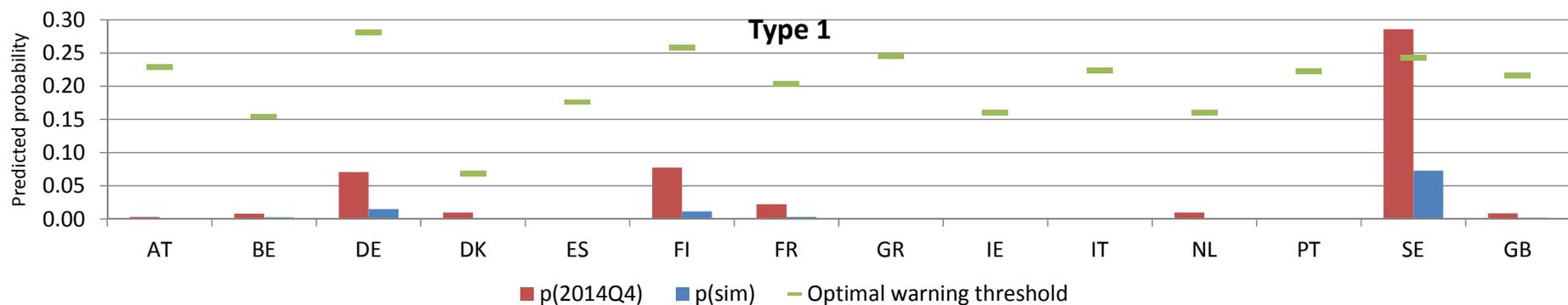
Impact on predicted probabilities I



Note: p(sim) corresponds to the average predicted probability over simulation horizon, assuming a capital ratio shock of 250 bps, where the optimal warning threshold corresponds to $\mu=0.85$.

- Cross-country heterogeneity: countries with predicted probabilities close to 0 in 2014Q4 do not have much to gain from capital-based measures

Impact on predicted probabilities II



Note: p(sim) corresponds to the average predicted probability over simulation horizon, assuming a capital ratio shock of 250 bps, where the optimal warning threshold corresponds to $\mu=0.85$.

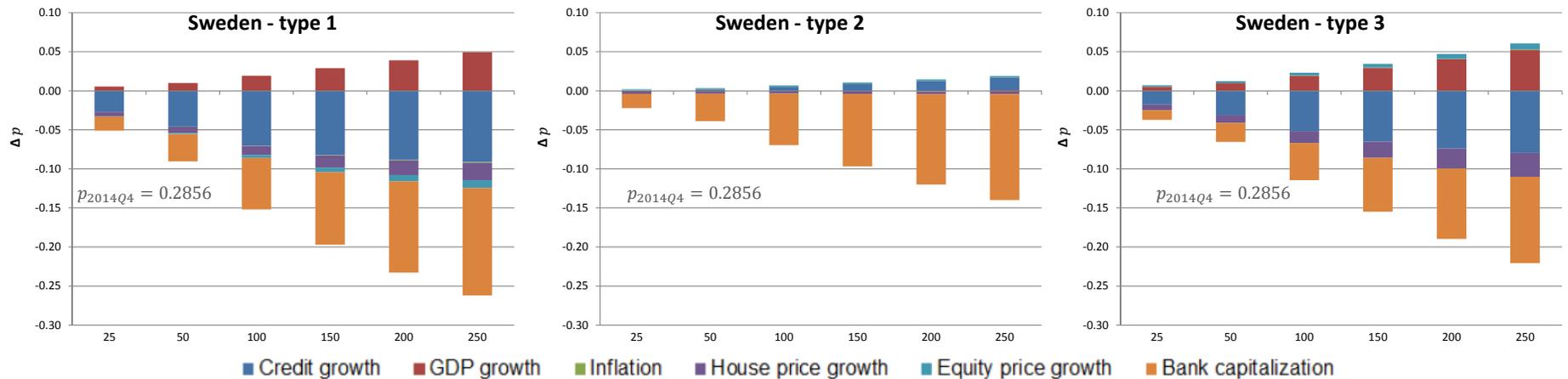
- Benchmark: bank credit model (Model 8, $\mu = 0.85$)
- Model is currently issuing a warning for Sweden
- Probability of being in a vulnerable state is reduced in all countries
- Cross-country heterogeneity: countries with predicted probabilities close to 0 in 2014Q4 do not have much to gain from capital-based measures

Impact on predicted probabilities III

| Type 1 shock | | | | | | |
|--------------------------------|-------|-------|-------|-------|-------|-------|
| Buffer size (in bps) | 25 | 50 | 100 | 150 | 200 | 250 |
| $-\Delta p$ in average country | 0.006 | 0.010 | 0.018 | 0.022 | 0.025 | 0.028 |
| ... relative to average p | 0.031 | 0.051 | 0.090 | 0.112 | 0.128 | 0.139 |
| ... relative to p in 2014Q4 | 0.176 | 0.295 | 0.517 | 0.649 | 0.739 | 0.802 |
| Type 2 shock | | | | | | |
| Buffer size (in bps) | 25 | 50 | 100 | 150 | 200 | 250 |
| $-\Delta p$ in average country | 0.003 | 0.004 | 0.008 | 0.011 | 0.014 | 0.016 |
| ... relative to average p | 0.013 | 0.022 | 0.042 | 0.056 | 0.068 | 0.078 |
| ... relative to p in 2014Q4 | 0.070 | 0.124 | 0.239 | 0.321 | 0.385 | 0.435 |
| Type 3 shock | | | | | | |
| Buffer size (in bps) | 25 | 50 | 100 | 150 | 200 | 250 |
| $-\Delta p$ in average country | 0.004 | 0.006 | 0.011 | 0.015 | 0.018 | 0.020 |
| ... relative to average p | 0.018 | 0.030 | 0.057 | 0.075 | 0.090 | 0.101 |
| ... relative to p in 2014Q4 | 0.109 | 0.191 | 0.360 | 0.478 | 0.569 | 0.640 |

- Reduction in crisis probabilities is greatest under type 1
- About half the size under type 2, somewhere in between under type 3

Transmission channels – contribution of variables



Note: The figures show the contribution of individual variables to the change in the probability of being in a vulnerable state relative to 2014Q4 for different capital ratio shocks.

- Type 1 & 3:

- Reduction in predicted probabilities mainly driven by increase in bank capitalization and reduction in credit growth; to a lesser extent: reduction in house price growth
- Reduction in GDP growth increases probability of being vulnerable

- Type 2:

- Reduction almost entirely driven by increase in banking sector capitalization
- Countervailing effects due to higher credit growth are of minor importance

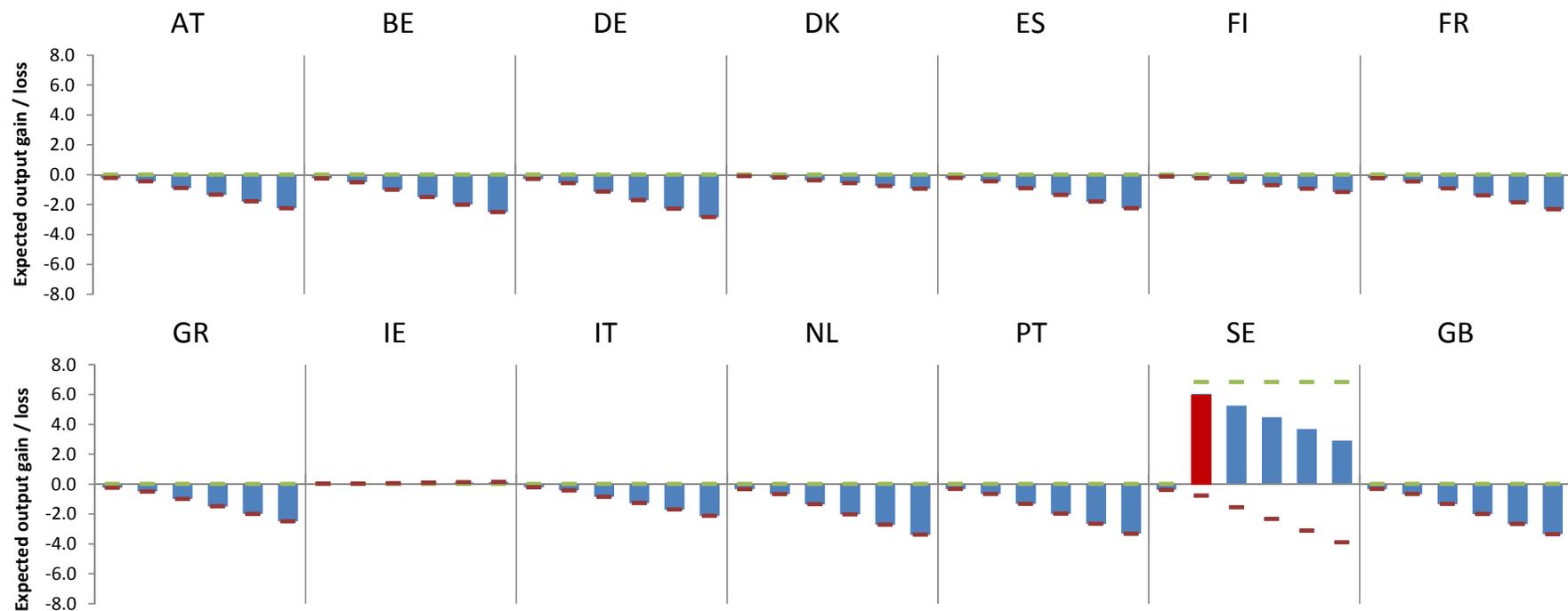
An alternative way to assess the net benefit

- Make stronger use of binary signals issued by the EWS
 1. Check whether the model is currently issuing a warning signal
 2. If yes, check whether capital ratio shock succeeds in pushing predicted probabilities below optimal early warning threshold
 3. If yes, calculate reduction in probability of being in vulnerable state, $-\Delta p$, as difference between probability of being in vulnerable state conditional on signal being issued and unconditional probability of being in vulnerable state
 4. Obtain estimate for benefit and net benefit as before (crisis cost unchanged):

$$\textit{benefit} = -\Delta p * \textit{crisis cost} \quad \text{and} \quad \textit{net benefit} = \textit{benefit} - \textit{cost}$$

- Benefit equal 0 if model currently not issuing warning signal or if the measure does not succeed in pushing probabilities below threshold
- Optimal buffer rate likely to be the lowest one that succeeds in pushing predicted probability below optimal threshold

Net benefits – alternative assessment (Type 3)



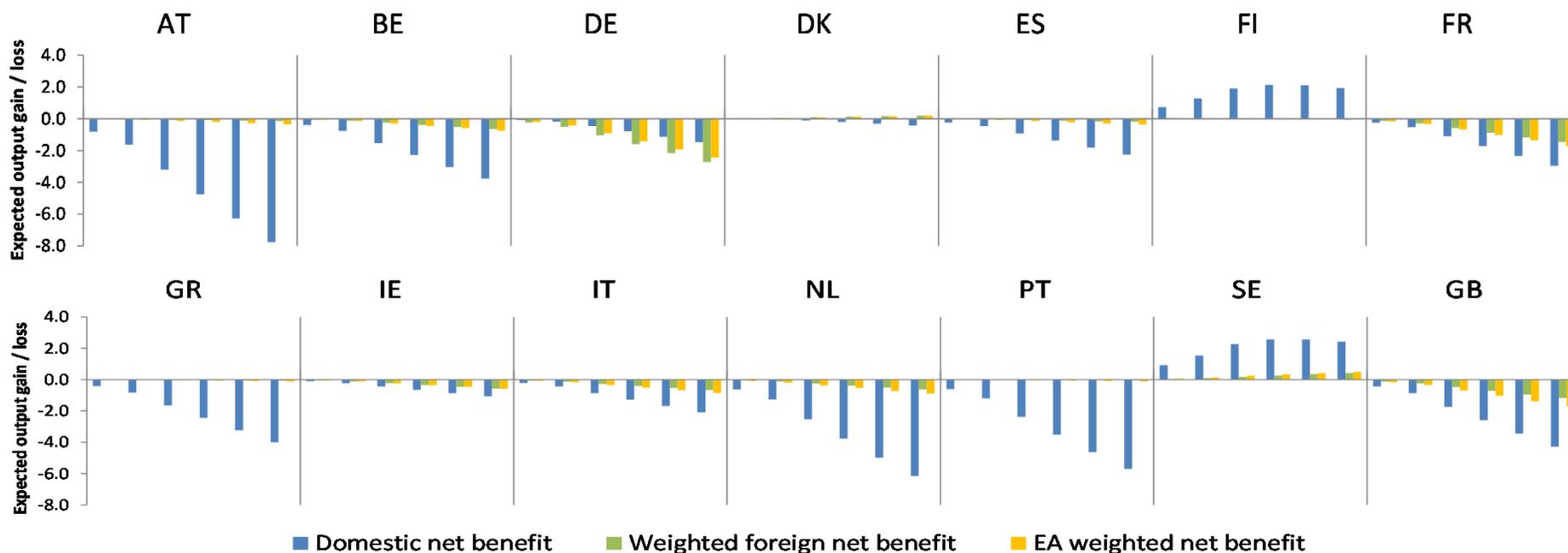
Note: Bars and markers for each country refer to capital ratio shocks of 25, 50, 100, 150, 200, and 250 bps, respectively.

- Sweden is the only country for which models are issuing a warning
- Capital ratio shock of 50 bps would be sufficient to push predicted probabilities below optimal warning threshold

Spillover effects

- GVAR model structure allows looking at spillover effects as it generates domestic and foreign responses
- Our model captures two important channels through which capital-based measures may exert spillover effects:
 - through banks that do significant cross-border lending
 - through the trade channel
- Calculate net benefits in both domestic and foreign countries
- Three variables:
 - **Domestic net benefit:** as before
 - **Weighted foreign net benefit:** net benefits in countries abroad, weighted by GDP as of 2014Q4 (alternatively: unweighted average)
 - **EA weighted net benefit:** net benefits in all sample countries, weighted by GDP in 2014Q4, no distinction between domestic and foreign effects

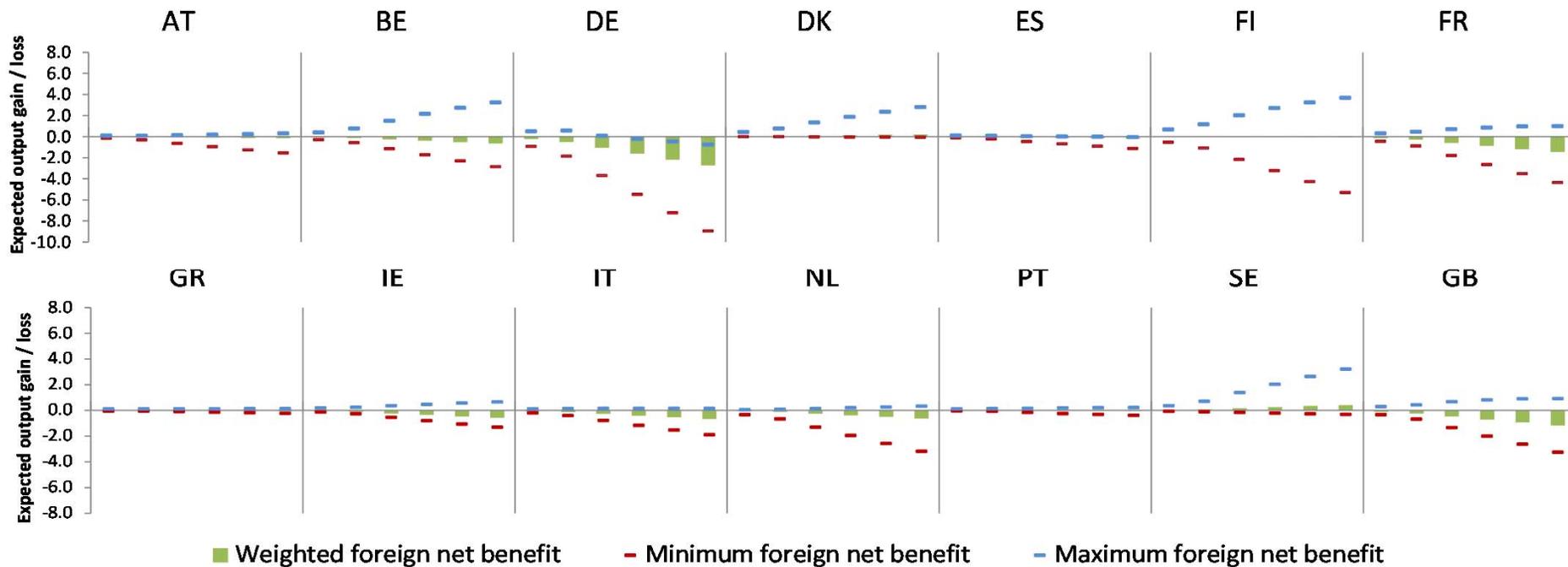
Spillovers – contractionary deleveraging (Type 1)



Note: Bars and markers for each country refer to capital ratio shocks of 25, 50, 100, 150, 200, and 250 bps, respectively.

- Significant spillover effects particularly for larger countries
- Aggregate foreign effects tend to go on in same direction as domestic effects

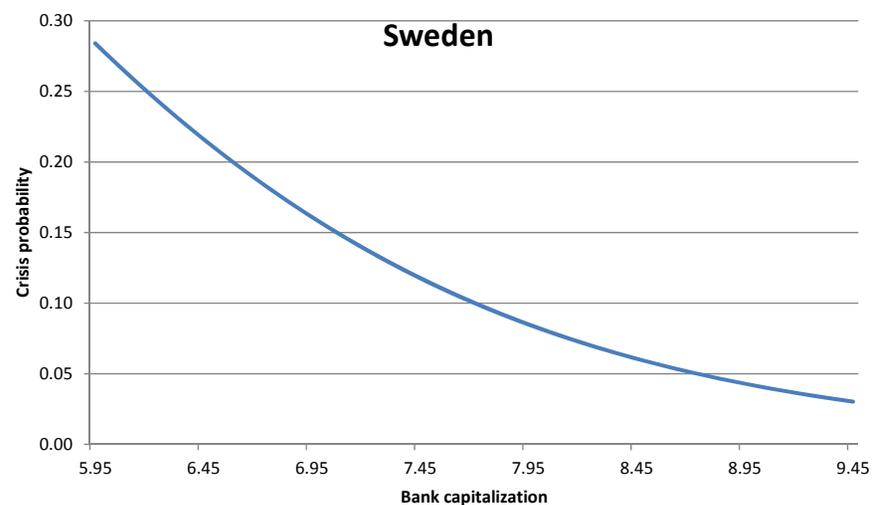
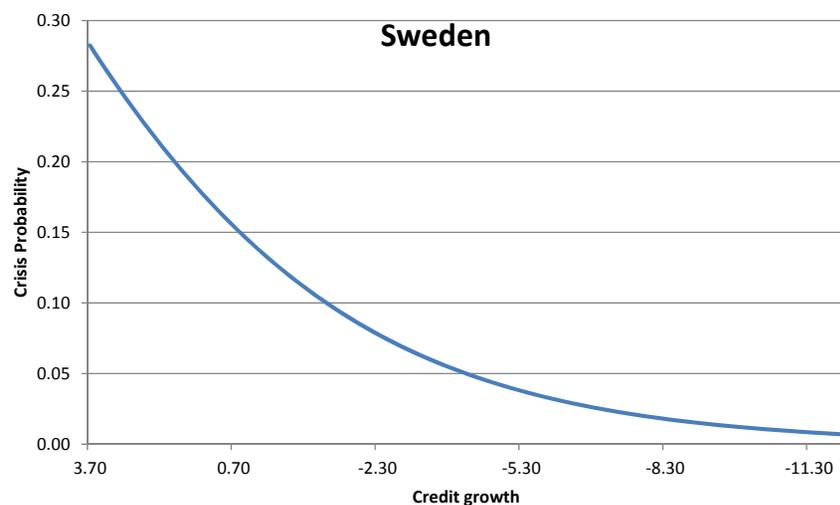
Heterogeneity in spillover effects (Type 1)



Note: Bars and markers for each country refer to capital ratio shocks of 25, 50, 100, 150, 200, and 250 bps, respectively.

- Spillover effects can go in different directions, considerable heterogeneity
- For larger countries (DE, ES, FR, IT, GB) they tend to be negative at the current stage of cycle

Nonlinearities in impact on predicted probabilities

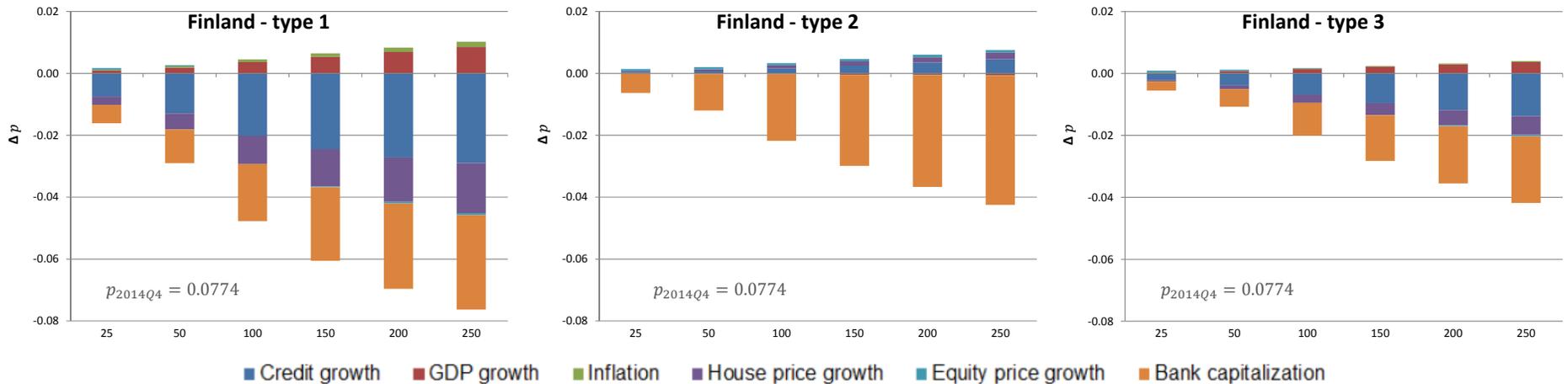


- Graphs show the influence of reductions in credit growth / increases in bank capitalization on predicted crisis probabilities, keeping all other variables fixed at their values in 2014Q4
- Convexity in the response functions: impact of further reductions / increases decreases in already achieved reductions / increases

Transmission channels – Calculation of contributions

- Aim: assess the relative contribution of variables in a nonlinear model
- We make use of the following 4-step procedure:
 1. Feed responses of individual variables (obtained from GVAR) into the logit model, **one by one**, assuming that the other variables remain at their 2014Q4 values, and obtain predicted probabilities
 2. For **each variable**, calculate the **change in the probability** of being in a vulnerable state as the difference between average predicted probability over the simulation horizon and the predicted probability in 2014Q4
 3. Sum the changes in the probability of being in a vulnerable state obtained from step 2, and calculate the **relative contribution of each variable** as the change in probability for that variable divided by the sum of all changes in probability
 4. Multiply the relative contribution of each variable from step 3 with the **aggregate reduction in crisis probabilities** for the respective country from the model
- Alternative: make use of marginal effects in 2014Q4 to obtain linear approximations for the effects of individual variables on crisis probabilities

Transmission channels – The case of Finland



Note: The figures show the contribution of individual variables to the change in the probability of being in a vulnerable state relative to 2014Q4 for different capital ratio shocks.

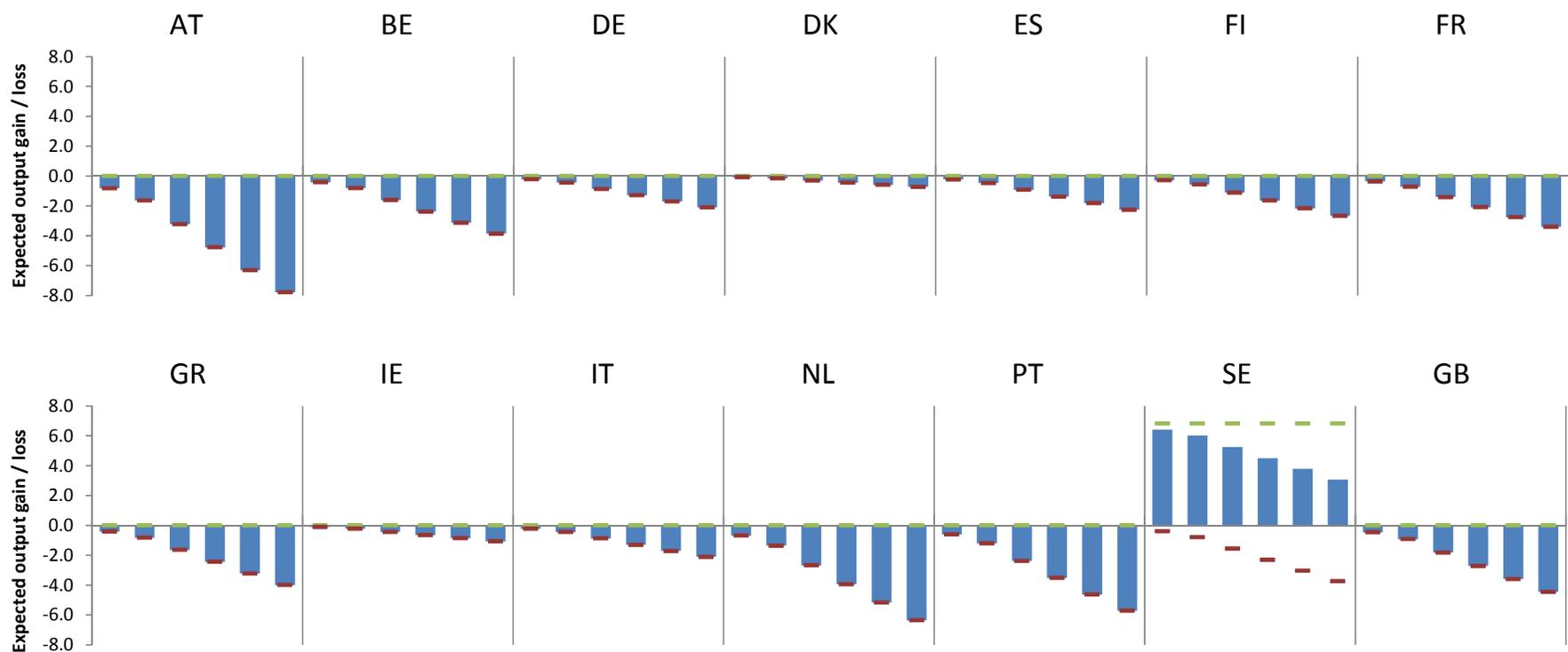
- Type 1 & 3:

- Reduction in predicted probabilities mainly driven by increase in bank capitalization and reduction in credit growth; to a lesser extent: reduction in house price growth
- Reduction in GDP growth increases probability of being vulnerable

- Type 2:

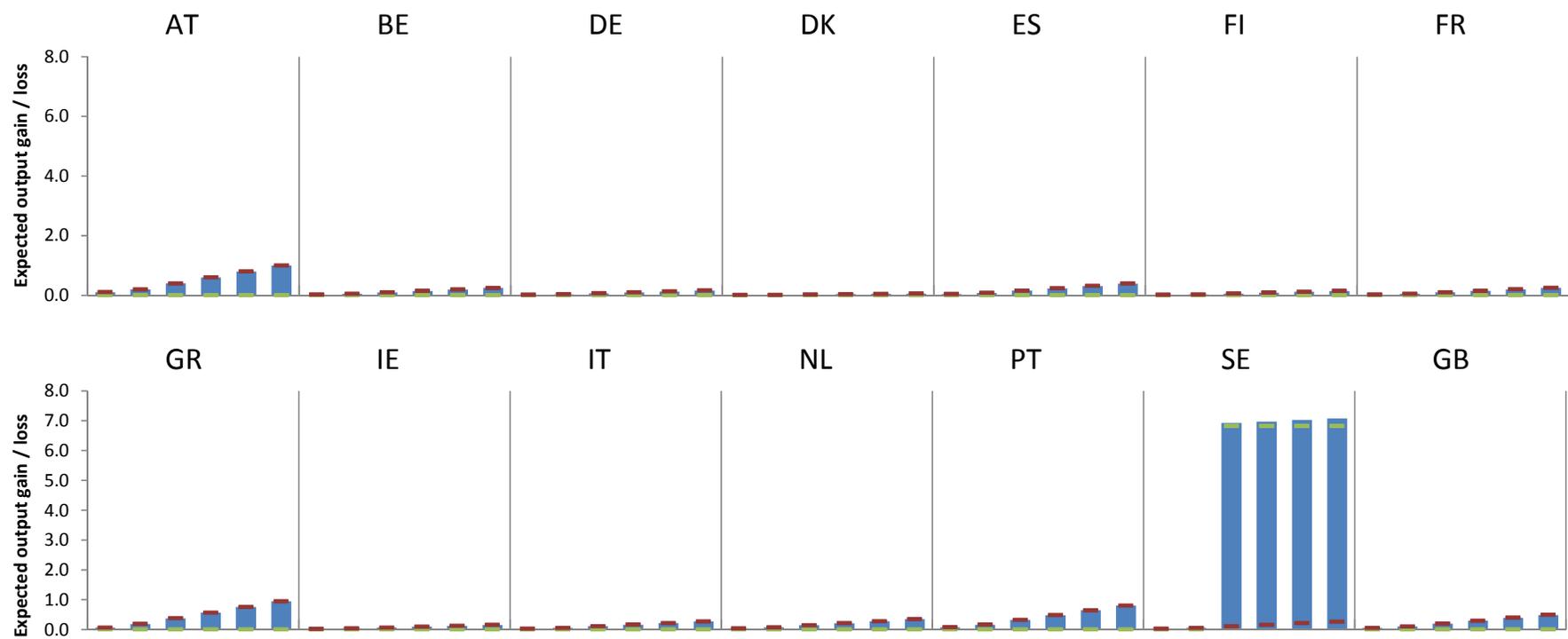
- Reduction almost entirely driven by increase in banking sector capitalization
- Countervailing effects due to higher credit growth are of minor importance

Net benefits – alternative assessment (Type 1)



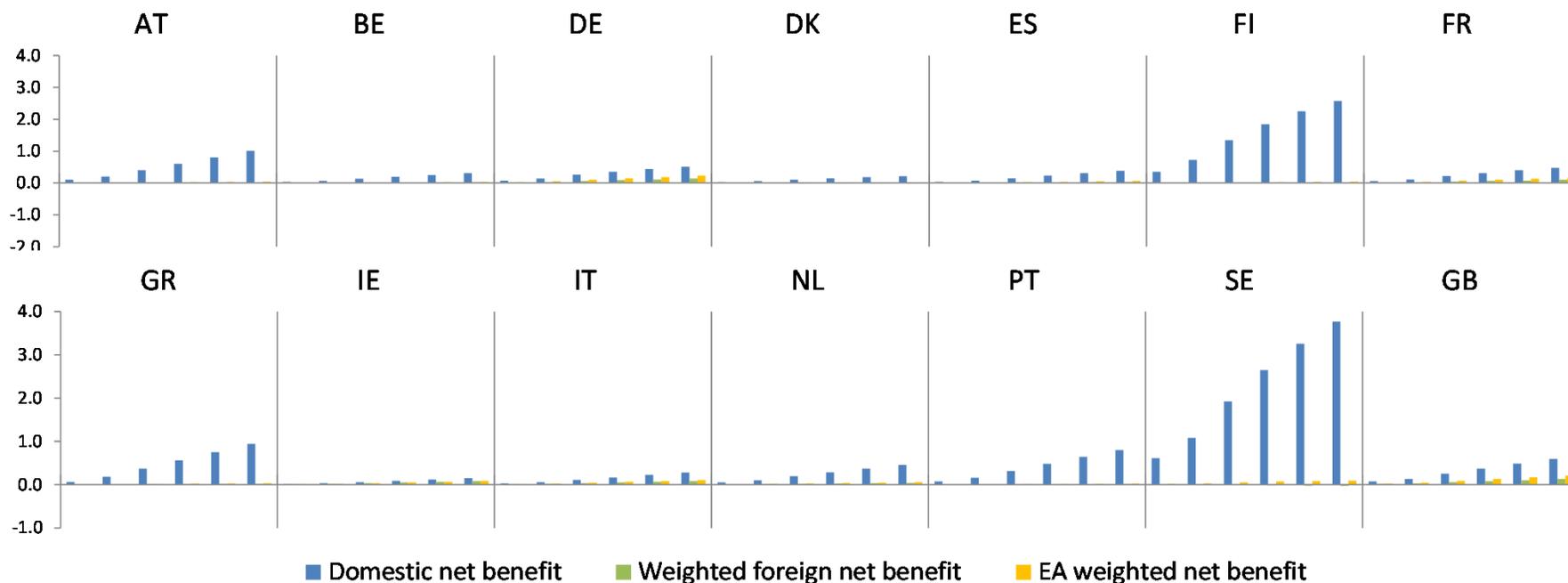
Note: Bars and markers for each country refer to capital ratio shocks of 25, 50, 100, 150, 200, and 250 bps, respectively.

Net benefits – alternative assessment (Type 2)



Note: Bars and markers for each country refer to capital ratio shocks of 25, 50, 100, 150, 200, and 250 bps, respectively.

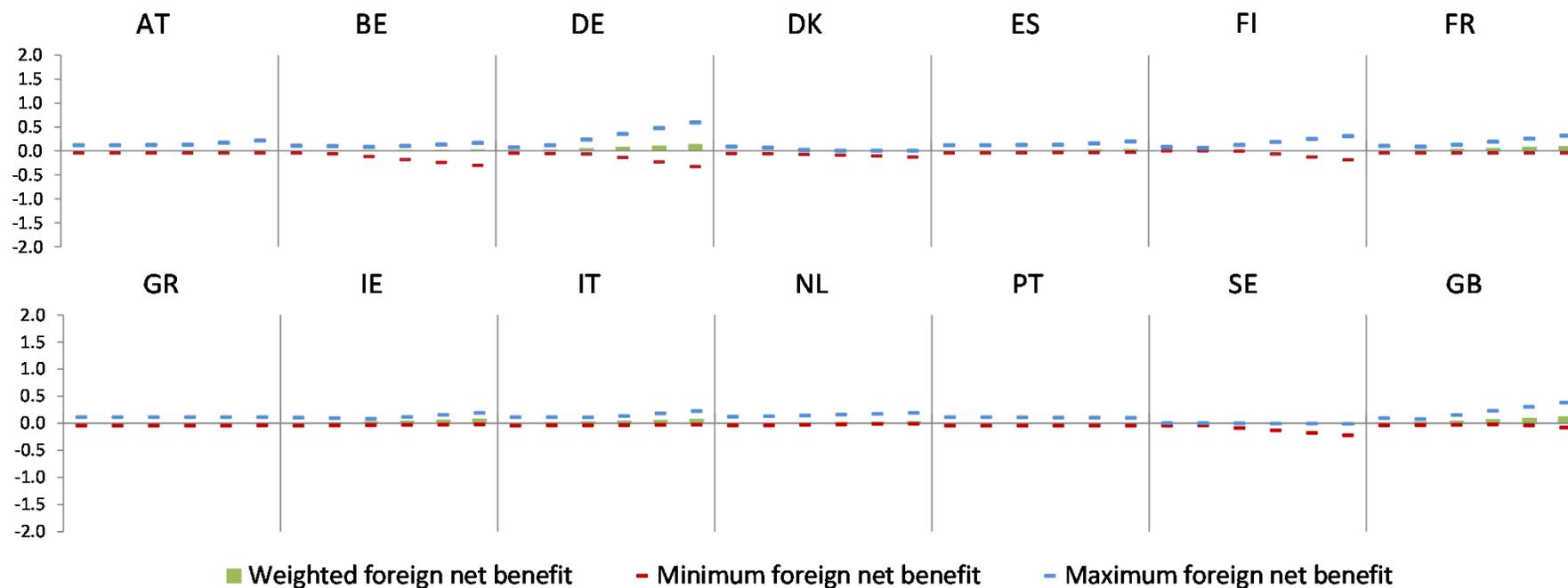
Spillovers – expansionary deleveraging (Type 2)



Note: Bars and markers for each country refer to capital ratio shocks of 25, 50, 100, 150, 200, and 250 bps, respectively.

- Spillover effects much smaller for type 2 shocks, since credit, GDP and asset price responses are much less pronounced

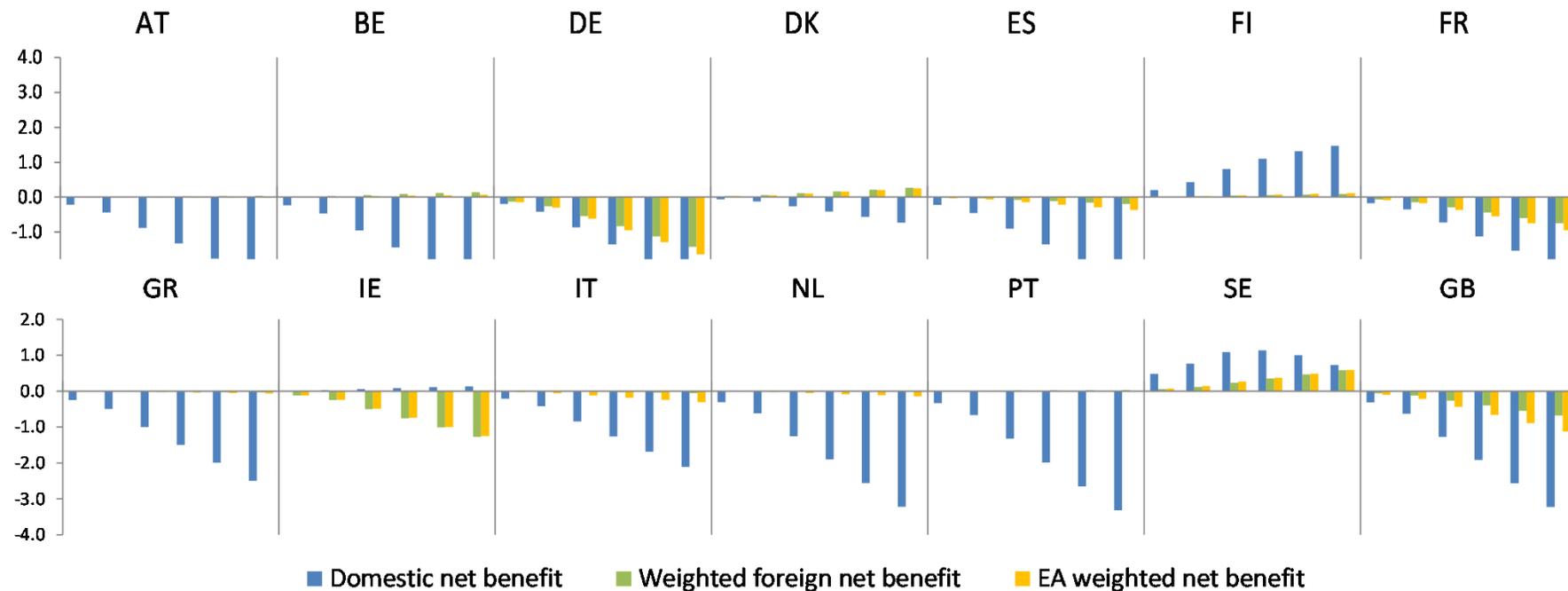
Heterogeneity in spillover effects (Type 2)



Note: Bars and markers for each country refer to capital ratio shocks of 25, 50, 100, 150, 200, and 250 bps, respectively.

- No significant spillover effects for type 2 shocks

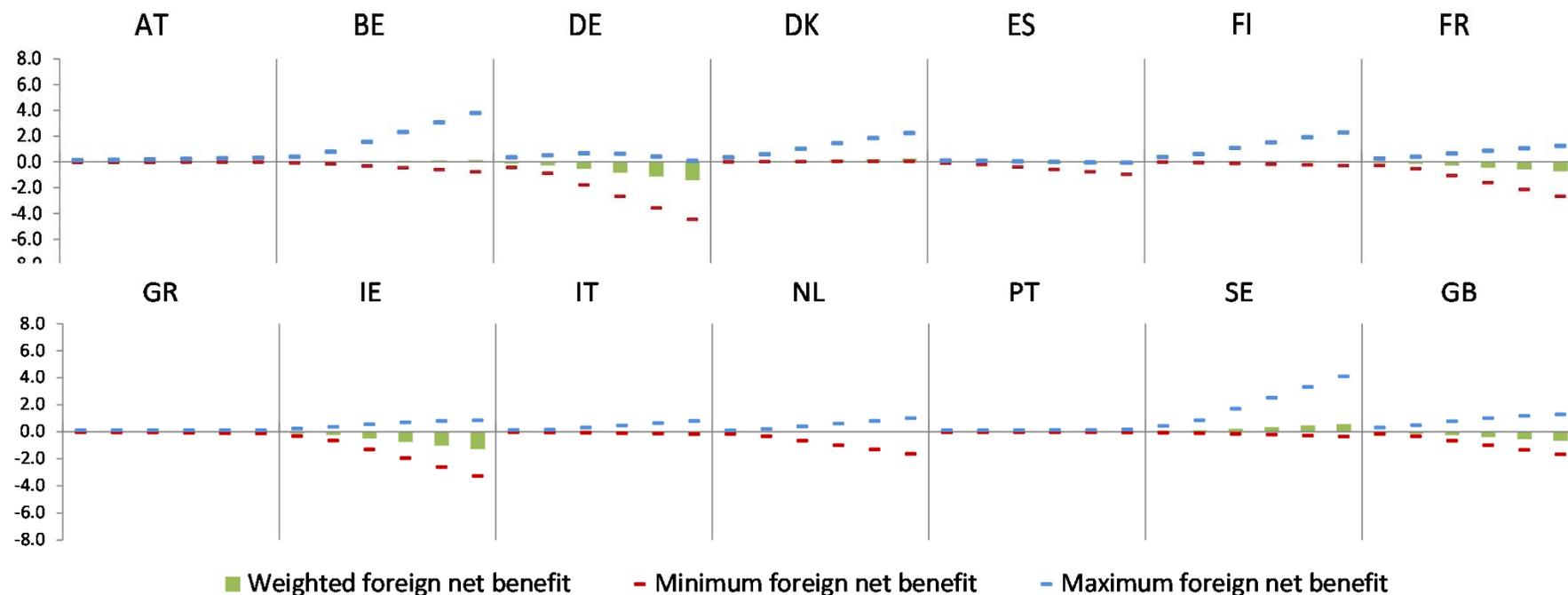
Spillovers – unconstrained deleveraging (Type 3)



Note: Bars and markers for each country refer to capital ratio shocks of 25, 50, 100, 150, 200, and 250 bps, respectively.

- As before, type 3 effects are somewhere in between, closer to type 1

Heterogeneity in spillover effects (Type 3)



Note: Bars and markers for each country refer to capital ratio shocks of 25, 50, 100, 150, 200, and 250 bps, respectively.

- Spillover effects under type 3 simulation somewhat smaller but resemble those under type 1

Models used to estimate long-term cost of new regulatory framework

Table A4.1

Key features of the models used in the analysis

| Model | Model type | Reference country/ area | Estimated/calibrated | Features bank capital | Features bank liquidity | Key lending spread ¹ |
|--------------------------------|-----------------|-------------------------|----------------------|-----------------------|-------------------------|---------------------------------|
| (1) Gerali et al (2010) | DSGE | euro area | largely estimated | yes | no | $i_l - i_d$ |
| (2) Roger and Vitek (2010) | DSGE | euro area | calibrated | yes | yes | $i_l - i_d$ |
| (3) Roeger ² (2010) | DSGE | euro area | calibrated | yes | yes | $i_l - i_d$ |
| (4) Christiano et al (2010) | DSGE | euro area | estimated | yes | yes | $i_l - i_d$ |
| (5) Antipa et al (2010) | DSGE | euro area | estimated | no | no | $i_l - i_d$ |
| (6) Roger and Vitek (2010) | DSGE | US | calibrated | yes | yes | $i_l - i_d$ |
| (7) Van den Heuvel (2008) | DGE | US | calibrated | yes | no | $i_l - i_d$ $i_e - i_d$ |
| (8) Curdia and Woodford (2009) | DSGE | US | estimated | no | no | $i_l - i_d$ |
| (9) Dellas et al (2010) | DSGE | US | calibrated | no | yes | $i_l - i_d$ |
| (10) Meh and Moran (2008) | DSGE | US | calibrated | yes | no | $i_l - i_d$ |
| (11) Locarno (2004) | Semi-structural | Italy | estimated | no | no | $i_l - i_d$ $i_b - i_d$ |
| (12) Bank of England | Semi-structural | UK | estimated | no | no | <i>n.a.</i> |
| (13) Gambacorta (2010) | VECM | US | estimated | yes | yes | $i_l - i_m$ |

¹ i_l : interest rate on loans to firms; i_b : interest rate on long-term bonds; i_d : interest rate on bank deposits; i_e : return on bank equity; i_m : monetary policy rate. ² Model calibrated based on eight euro area countries.

Source: BIS 2010, An assessment of the long-term economic impact of stronger capital and liquidity requirements

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